

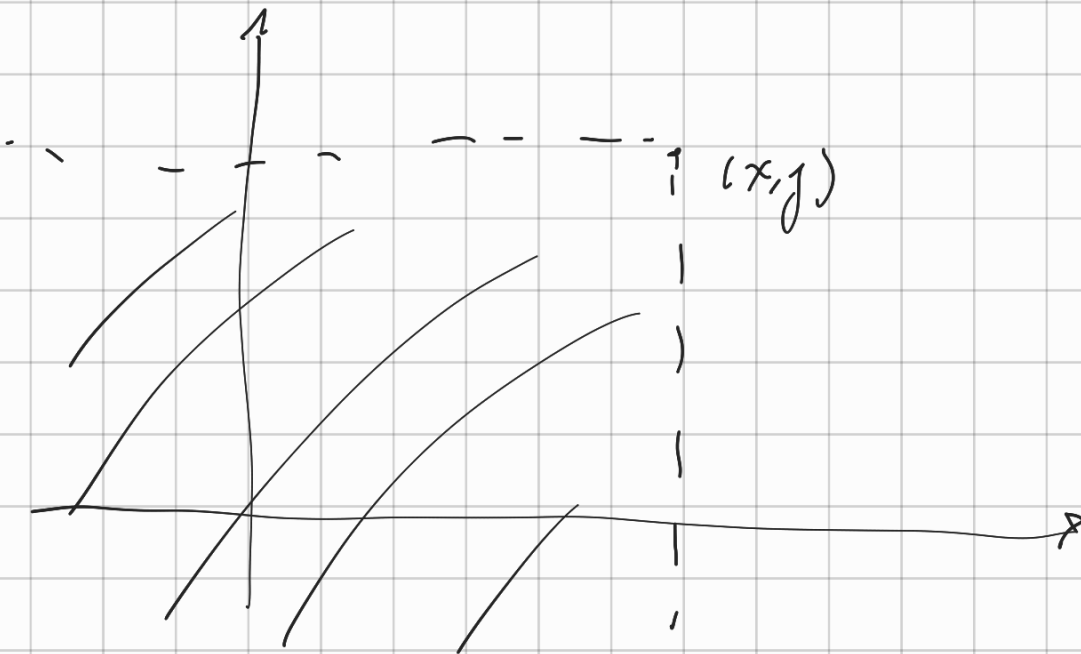
# Probability Theory

10/13/22

Jointly distributed cont. r.v.

$$F(x) = \mathbb{P}(X \leq x) \quad \text{I r.v.}$$

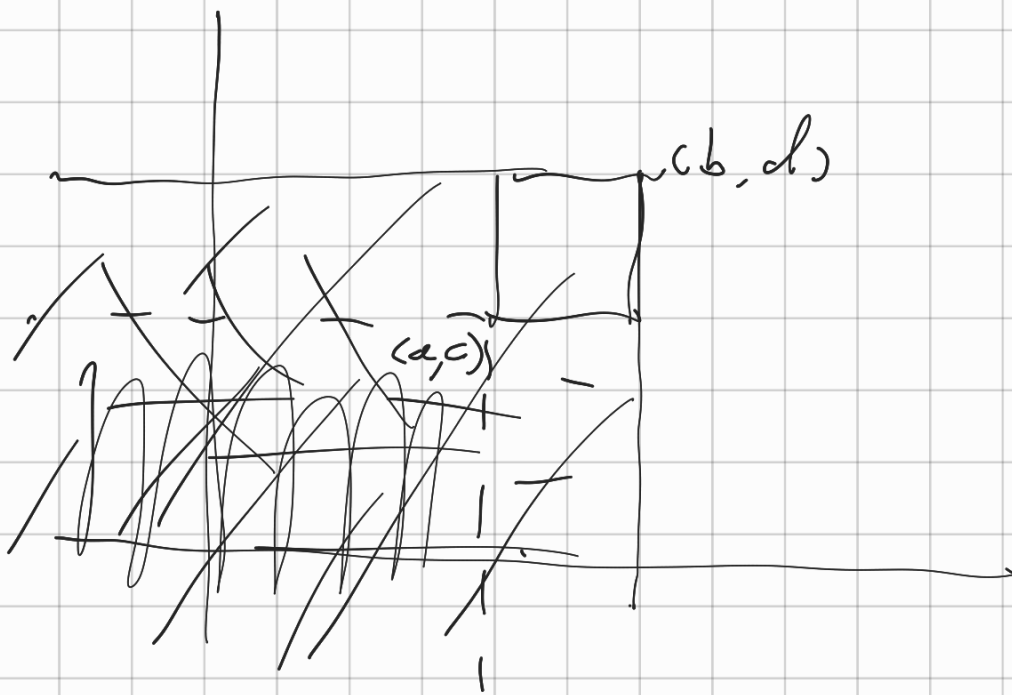
$$F(x, y) = \mathbb{P}(X \leq x \ \& \ Y \leq y)$$



$F(x, y)$

$$\mathbb{P}(a < X \leq b \text{ and } c < Y \leq d) =$$

$$F(b, d) - F(b, c) - F(a, d) + F(a, c)$$



Once you have the prob. of rectangles,  
 $X$  and  $Y$  are jointly continuous  
 if  $F(x, y)$  is differentiable.

$$P(x < X < x + dx, y < Y < y + dy) = \int_x^{x+dx} \int_y^{y+dy} f(x, y) dx dy$$

$$f(x, y) = \frac{\partial^2}{\partial x \partial y} F(x, y)$$

$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(u, v) du dv$$

Uniform dist.

$X, Y$  are uniform in

$$[0, 1] \times [0, 1]$$

$$f(x, y) = \begin{cases} 1 & \text{if } 0 \leq x, y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$f(x, y) = f_X(x) f_Y(y)$$

$$f_X(x) = \begin{cases} 1 & \text{if } 0 \leq x \leq 1 \end{cases}$$

$X$  and  $Y$  are independent.

Uniform in  $[A, B] \times [C, D]$

$$f(x, y) = \frac{1}{(B-A) \cdot (D-C)} \quad \begin{array}{l} A \leq x \leq B \\ C \leq y \leq D \end{array}$$

if  $A$  is a subset of  $\mathbb{R}^2$  (e.g. a circle)

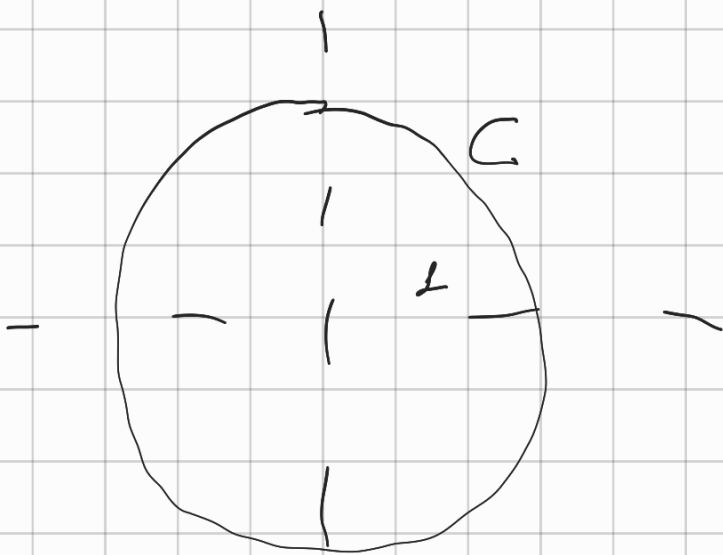
uniform in  $A$

$$f(x, y) = \begin{cases} \frac{1}{|A|} & \text{if } (x, y) \in A \\ 0 & \text{otherwise} \end{cases}$$

In The case of rectangle

$$f(x, y) = f_x(x) f_y(y) \Rightarrow \text{indep.}$$

in The case of circle  $\Rightarrow$  not indep.



$$|C| = \pi$$

$$f(x, y) = \begin{cases} c & \text{if } (x, y) \in C \\ 0 & \text{otherwise} \end{cases}$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1 = c |C|$$

## Rectangle

$$f(x, y) = g(x) h(y)$$

$$g(x) = \frac{1}{(B-A)(C-D)}$$

$$A \leq x \leq B$$

$$h(y) = 1$$

$$C \leq y \leq D$$

□

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## Marginals

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

Uniform in

$[A, B] \times [C, D]$

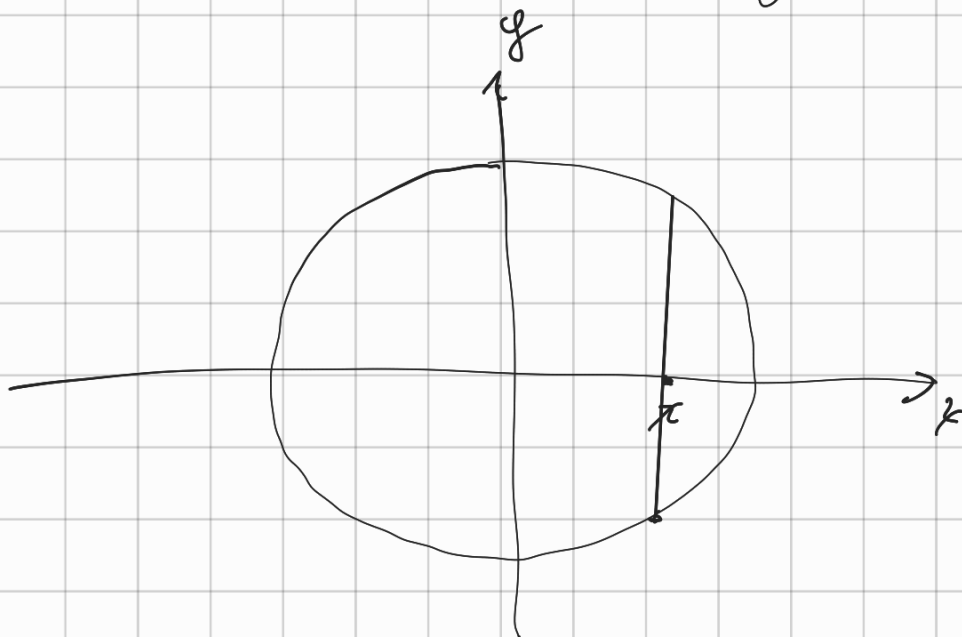
$$f_X(x) = \frac{1}{B-A}$$

$$A \leq x \leq B$$

$$f_Y(y) = \frac{1}{D-C}$$

$$C \leq y \leq D$$

Uniform on a circle of radius  
1 cent. at the origin



$$f_x(x) = 2\sqrt{1-x^2} \quad -1 \leq x \leq 1$$

$$f_y(y) = 2\sqrt{1-y^2} \quad -1 \leq y \leq 1$$

$$f(x, y) \neq f_x(x) f_y(y) !$$

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$$f(x, y) = \frac{1}{\pi}$$

$$(x, y) \in C$$

In dependence

$X \perp\!\!\!\perp Y$  if

$$f(x, y) = f_X(x) f_Y(y)$$

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Exponential r.v.

$$X \quad f_X = \lambda e^{-\lambda x} \quad x > 0$$

$X \quad Y$  independent

$$Z = X + Y$$

$$P(Z \leq z) = P(X + Y \leq z) =$$

$$= \int_A f(x, y) dx dy =$$

$$A = \{ (x, y) \mid x + y \leq z \}$$

$$= \int_{x=-\infty}^{\infty} \int_{y=-\infty}^{z-x} f(x, y) dx dy =$$

$$= F(z)$$

$$f_z(z) = \frac{d}{dz} F(z) = \int_{-\infty}^{\infty} f(x, z-x) dx$$

if  $X$   $Y$  are indep.

$$f_z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx$$

$$f_z = f_X * f_Y$$

$$f_X(x) = \lambda e^{-\lambda x} \quad x > 0$$

$$f_Y(y) = \mu e^{-\mu y} \quad y > 0$$



$$f_Z(z) = \lambda \mu \int_0^z e^{-\lambda x} e^{-\mu(z-x)} dx$$

$$\lambda = \mu$$

$$= \lambda \int_0^z e^{-\lambda x - \lambda(z-x)} dx =$$

$$= \lambda z e^{-\lambda z}$$

$$X_3 \quad \lambda e^{-\lambda x}$$

$$Z + X_3 = Z_3$$

$$f_{Z_3}(z) = \lambda \int_0^z \underbrace{x e^{-\lambda x}}_Z \underbrace{e^{-\lambda(z-x)}}_{X_3} dx$$

$$= \lambda^2 e^{-\lambda z} \int_0^z x dx = \lambda^2 e^{-\lambda z} \frac{z^2}{2}$$

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If I have  $n$  exponential v.v.

$X_1, X_2, \dots, X_n$  i.i.d.

$$Z_n = X_1 + X_2 + \dots + X_n$$

$$f_n(z) = \frac{\lambda^n z^{n-1}}{(n-1)!} e^{-\lambda z}$$

Prob that before Time  $t$  I get  
 $n$  events.

$$P\left(\sum_{i=1}^n X_i \leq t \ \& \ \sum_{i=1}^{n+1} X_i > t\right) =$$
$$\frac{(\lambda t)^n}{n!} e^{-\lambda t}$$

Poissonian with par  $\lambda t$ .

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Conditional distribution

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$$

$$f_{Y|X}(y|x) = \frac{f(x,y)}{f_X(x)}$$

$$X \perp\!\!\!\perp Y \iff f_{Y|X}(y|x) = f_Y(y).$$